# Invariant Risk Minimization

### Authors: Martin Arjovsky, L´eon Bottou, Ishaan Gulrajani, David Lopez-Paz

February 21, 2025

Authors: Martin Arjovsky, L´eon Bottou, Isha

Invariant Risk Minimization

February 21, 2025

### ML training is done via minimizing some training loss





#### Figure: Task: Classification of cows vs camels

# Motivation

#### The problem



(a) Grassy background



#### (b) Sandy background

Figure: Training data contains biases



Camel?

Authors: Martin Arjovsky, L´eon Bottou, Isha

Invariant Risk Minimization

**Correlations-vs-causations** Minimizing training error leads machines into recklessly absorbing all the correlations found in training data. Spurious correlations (landscape, contexts) are unrelated to causal explanations of interest (animal shapes) **Causation** Correlations that are stable (invariant) across training environments.

**Invariant Risk Minimization (IRM) principle** To learn invariances across environments, find a data representation such that the optimal classifier on top of that representation matches for all environments.

1. IRM training objective to learn invariance features across different **training** environments

2. After achieving the desired invariance and a model with low error across training environments, we want to know:

a. When do these conditions imply invariance across all environments

b. When do these conditions lead to low error across **all** environments (basically, OOD generalization)

c. Connect invariance and OOD generalization to theory of causation

Datasets  $D_e := \{(x_i^e, y_i^e)_{i=1}^{n_e} \text{ under multiple environments } e \in \mathcal{E}_{tr}$ A large set of unseen but related environments  $\mathcal{E}_{all} \supset \mathcal{E}_{tr}$ Intuitive goal: Learn predictor  $Y \approx f(X)$  that performs well across  $\mathcal{E}_{all}$ Denote

$$R^{e}(f) := \mathbb{E}_{X^{e}, Y^{e}}[\ell(f(X^{e}), Y^{e})]$$

is risk under environment e

Take  $\ell = MSE$  or cross-entropy, then the *optimal predictors* can be written as conditional expectations.

Take  $\ell = MSE$  or cross-entropy, then the *optimal predictors* can be written as conditional expectations.

We say a data representation  $\Phi : \mathcal{X} \to \mathcal{H}$  elicits an invariant predictor across environment  $\mathcal{E}$  if and only if

$$\mathbb{E}[Y^e \mid \Phi(X^e) = h] = \mathbb{E}[Y^{e'} \mid \Phi(X^{e'}) = h]$$

 $\forall h \in \cap_{e \in \mathcal{E}} \operatorname{supp} (\Phi(X^e))$ 

Take  $\ell = MSE$  or cross-entropy, then the *optimal predictors* can be written as conditional expectations.

We say a data representation  $\Phi : \mathcal{X} \to \mathcal{H}$  elicits an invariant predictor across environment  $\mathcal{E}$  if and only if

$$\mathbb{E}[Y^e \mid \Phi(X^e) = h] = \mathbb{E}[Y^{e'} \mid \Phi(X^{e'}) = h]$$

 $\forall h \in \cap_{e \in \mathcal{E}} \operatorname{supp} (\Phi(X^e))$ 

**Formal Def** Say data representation  $\Phi$  elicits an invariant predictor  $w \circ \Phi$  across  $\mathcal{E}$  if there is a classifier  $w : \mathcal{H} \to \mathcal{Y}$  simultaneously optimal  $\forall e \in \mathcal{E}$ :

 $w \in \arg\min_{\bar{w}} R^e(\bar{w} \circ \Phi)$  (optimization constraint)

$$\begin{split} & \min_{\substack{\Phi:\mathcal{X}\to\mathcal{H}\\w:\mathcal{H}\to\mathcal{Y}}} & \sum_{e\in\mathcal{E}_{\mathrm{tr}}} R^e(w\circ\Phi) \\ & \text{subject to} & w\in\arg\min_{\bar{w}:\mathcal{H}\to\mathcal{Y}} R^e(\bar{w}\circ\Phi), \text{ for all } e\in\mathcal{E}_{\mathrm{tr}}. \end{split}$$
(IRM)

Instantiate IRM into the practical version (derived in the paper):

$$\min_{\Phi:\mathcal{X}\to\mathcal{Y}}\sum_{e\in\mathcal{E}_{tr}}R^{e}(\Phi)+\lambda\cdot\|\nabla_{w|w=1.0}R^{e}(w\cdot\Phi)\|^{2}, \quad (\mathsf{IRMv1})$$

w = 1 is a scalar and fixed "dummy" classifier,  $\lambda \in [0, \infty)$  is a regularizer balancing between predictive power and the invariance of the predictor  $1 \cdot \Phi$ 

Estimate the objective IRMv1 using mini-batches for stochastic gradient descent (unbiased),

$$\sum_{k=1}^{b} \left[ \nabla_{w|w=1.0} \ell(w \cdot \Phi(X_k^{e,i}), Y_k^{e,i}) \cdot \nabla_{w|w=1.0} \ell(w \cdot \Phi(X_k^{e,j}), Y_k^{e,j}) \right],$$

where  $(X^{e,i}, Y^{e,i})$  and  $(X^{e,j}, Y^{e,j})$  are two random mini-batches of size *b* from environment *e*.

#### 1. Phrasing the constraints as a penalty

$$L_{\rm IRM}(\Phi, w) = \sum_{e \in \mathcal{E}_{\rm tr}} R^e(w \circ \Phi) + \lambda \cdot \mathbb{D}(w, \Phi, e) \tag{1}$$

 $\mathbb{D}(w, \Phi, e)$  measures how close w is to minimizing  $R^e(w \circ \Phi)$ , and  $\lambda \in [0, \infty)$  is a hyper-parameter balancing predictive power and invariance.

# Going from IRM to IRMv1

### 2. Choosing a penalty $\mathbb D$ for linear classifiers w

Consider learning an invariant predictor  $w \circ \Phi$ , where w is a linear-least squares regression, and  $\Phi$  is a nonlinear data representation.



Figure: Different measures of invariance lead to different optimization landscapes. The naïve approach of measuring the distance between optimal classifiers  $\mathbb{D}_{\rm dist}$  leads to a discontinuous penalty (solid blue unregularized, dashed orange regularized). In contrast, the penalty  $\mathbb{D}_{\rm lin}$  does not exhibit these problems.

#### 3. Fixing the linear classifier w

We recognize that when optimizing over  $(\Phi, w)$  using  $\mathbb{D}_{\text{lin}}$ , a pair  $(\gamma \Phi, \frac{1}{\gamma}w)$  can pick  $\gamma \approx 0$  to drive  $\mathbb{D}_{\text{lin}}$  towards zero without touching the risk term. Similarly, note:

$$w \circ \Phi = \underbrace{\left(w \circ \Psi^{-1}
ight)}_{\widetilde{w}} \circ \underbrace{\left(\Psi \circ \Phi
ight)}_{\widetilde{\Phi}}.$$

 $\rightarrow$  Can always re-parametrize our invariant predictor w and restrict it to be some non-zero value  $\tilde{w}$  of our choosing. This turns (1) into a relaxed version of IRM, where optimization only happens over  $\Phi$ :

$$\mathcal{L}_{\mathrm{IRM},w=\tilde{w}}(\Phi) = \sum_{e \in \mathcal{E}_{\mathrm{tr}}} R^{e}(\tilde{w} \circ \Phi) + \lambda \cdot \mathbb{D}_{\mathrm{lin}}(\tilde{w}, \Phi, e).$$
(2)

# Going from IRM to IRMv1

#### Scalar fixed classifiers $\tilde{w}$ are sufficient to monitor invariance

#### Theorem

For all  $e \in \mathcal{E}$ , let  $R^e : \mathbb{R}^d \to \mathcal{R}$  be convex differentiable cost functions. A vector  $v \in \mathbb{R}^d$  can be written  $v = \Phi^\top w$ , where  $\Phi \in \mathbb{R}^{p \times d}$ , and where  $w \in \mathbb{R}^p$  simultaneously minimize  $R^e(w \circ \Phi)$  for all  $e \in \mathcal{E}$ , if and only if  $v^\top \nabla R^e(v) = 0$  for all  $e \in \mathcal{E}$ . Furthermore, the matrices  $\Phi$  for which such a decomposition exists are the matrices whose nullspace  $\operatorname{Ker}(\Phi)$  is orthogonal to v and contains all the  $\nabla R^e(v)$ .

 $\rightarrow$  Any linear invariant predictor can be decomposed as linear data representations of different ranks.

 $\rightarrow$  can restrict our search to matrices  $\Phi \in \mathbb{R}^{1 \times d}$  and let  $\tilde{w} \in \mathbb{R}^1$  be the fixed scalar 1.0. This translates (2) into:

$$L_{\text{IRM},w=1.0}(\Phi^{\top}) = \sum_{e \in \mathcal{E}_{\text{train}}} R^{e}(\Phi^{\top}) + \lambda \cdot \mathbb{D}_{\text{lin}}(1.0, \Phi^{\top}, e).$$
(3)

IRM: promotes low error and invariance across training environments  $\mathcal{E}_{tr}$ 

 $\stackrel{?}{\rightarrow} \mathsf{Invariance} + \mathsf{low} \; \mathsf{error} \; \mathsf{across} \; \mathcal{E}_{\mathsf{all}}$ 

Invariance  $\stackrel{?}{\leftrightarrow}$  causality  $\stackrel{?}{\leftrightarrow}$  OOD generalization

**1.** Environments  $\circ$  The data from all the environments share the same underlying Structural Equation Model C := (S, N) over the feature and outcome vector  $(X_1, \ldots, X_d, Y)$ 

$$\mathcal{S}: X_i \leftarrow f_i(\mathsf{PA}(X_i), N_i)$$

• Then  $\mathcal{E}_{all}(\mathcal{C})$  indexes all the interventional distributions  $P(X^e, Y^e)$  obtainable by valid interventions e

**1.** Environments  $\circ$  The data from all the environments share the same underlying Structural Equation Model C := (S, N) over the feature and outcome vector  $(X_1, \ldots, X_d, Y)$ 

 $S: X_i \leftarrow f_i(\mathsf{PA}(X_i), N_i)$ 

 $\circ$  Then  $\mathcal{E}_{\mathsf{all}}(\mathcal{C})$  indexes all the interventional distributions  $P(X^e,Y^e)$  obtainable by valid interventions e

 $\circ$  Intervention *e* is valid if they "do not destroy too much information about the target variable *Y*":

The causal graph remains acyclic,

$$\mathbb{E}[Y^e \mid \mathsf{Pa}(Y)] = \mathbb{E}[Y \mid \mathsf{Pa}(Y)],$$

 $\mathbb{V}[Y^e | Pa(Y)]$  remains within a finite range.

Invariance  $\leftrightarrow$  Causation: predictor  $v : \mathcal{X} \to \mathcal{Y}$  is invariant on  $\mathcal{E}_{all} \Leftrightarrow$ attains optimal  $R^{OOD} \Leftrightarrow$  uses only the direct causal parents of Y to predict,  $v(x) = \mathbb{E}_{N_Y}[f_Y(Pa(Y), N_Y)]$ 

$$R^{\text{OOD}} = \max_{e \in \mathcal{E}_{\text{all}}} R^e(f)$$

 $\circ$  Diversity requirement: limits the extent to which the training environments are co-linear

#### Assumption

A set of training environments  $\mathcal{E}_{tr}$  lie in linear general position of degree r if  $|\mathcal{E}_{tr}| > d - r + \frac{d}{r}$  for some  $r \in$ , and for all non-zero  $x \in d$ :

$$\dim\left(\operatorname{span}\left(\left\{X^{e}\left[X^{e}X^{e}^{\top}\right]x-_{X^{e},\epsilon^{e}}\left[X^{e}\epsilon^{e}\right]\right\}_{e\in\mathcal{E}_{tr}}\right)\right)>d-r.$$

### 2. Invariant Causal Prediction (ICP) theory (Peters, 2015)

Theorem (Invariant Causal Prediction - ICP)

Consider a (linear) Gaussian SEM with interventions. Then given the identifiable causal predictors  $S(\mathcal{E})$  under interventions  $\mathcal{E}$ , all causal predictors are identifiable, that is

 $S(\mathcal{E}) = Pa(Y)$ 

*if the interventions are do-interventions, noise interventions or simultaneous noise interventions* 

 $\rightarrow$  IRM allows for non-Gaussian data, for linear transformation of the variables with stable and spurious correlations, does not require specific types of interventions or the existence of a causal graph

2. Invariant Causal Prediction (ICP) theory (Peters, 2015) **Theorem** (roughly stated): If one finds a representation  $\Phi \in \mathbb{R}^{d \times d}$  of rank r eliciting an invariant predictor  $w \circ \Phi$  across  $\mathcal{E}_{tr}$ , and  $\mathcal{E}_{tr}$  satisfying the diversity requirement, then  $w \circ \Phi$  is invariant across  $\mathcal{E}_{all}$ .

The setting in consideration: •  $Y^e = Z_1^e \cdot \gamma + \epsilon^e$ ,  $Z_1^e \perp \epsilon^e$ ,  $\mathbb{E}[\epsilon^e] = 0$ .  $Z_1$ : causal variables,  $Z_2$ : non-causal variables

•  $X^e = S(Z_1^e, Z_2^e)$ ,  $Z_1$  component of S is invertible

3. OOD generalization (low error) across  $\mathcal{E}_{tr}$  + invariance across  $\mathcal{E}_{all} = \mathbf{OOD}$  generalization across  $\mathcal{E}_{all}$ 

$$\Rightarrow$$
 Invariance  $\leftrightarrow$  OOD generalization

# Experiments results

Synthetic data generation process.



Figure 3: In our synthetic experiments, the task is to predict  $Y^e$  from  $X^e = S(Z_1^e, Z_2^e)$ .

Along with the following variations

◦ Scrambled (S) observations, where S is an orthogonal matrix, or unscrambled (U) observations, where S = I. ◦ Fully-observed (F) graphs, where  $W_{h\rightarrow 1} = W_{h\rightarrow y} = W_{h\rightarrow 2} = 0$ , or partially-observed (P) graphs, where  $(W_{h\rightarrow 1}, W_{h\rightarrow y}, W_{h\rightarrow 2})$  are Gaussian. ◦ Homoskedastic (O) Y-noise, where  $\sigma_y^2 = e^2$  and  $\sigma_2^2 = 1$ , or heteroskedastic (E) Y-noise, where  $\sigma_y^2 = 1$  and  $\sigma_2^2 = e^2$ . ◦ The 3 training environments are  $e \in \{0.2, 2, 5\}$  and we draw 1000 samples from each environment.

## Experiments results



Figure 4: Average errors on causal (plain bars) and non-causal (striped bars) weights for our synthetic experiments. The y-axes are in log-scale. See main text for details.

∃ →

Color each image in MNIST with either red or green in a way that correlates strongly (but spuriously) with the class label.

Three environments (two training, one test) formed by:

• Assign a preliminary binary label  $\tilde{y}$  based on the digit:  $\tilde{y} = 0$  for digits 0-4 and  $\tilde{y} = 1$  for digits 5-9, then flip  $\tilde{y}$  with probability 0.25 to get the final label y.

• Sample a color ID z by flipping y with probability  $p_e$ , which is 0.2 (first environment), 0.1 (second), or 0.9 (test).

 $\circ$  Color each image red if z = 1 or green if z = 0.

## Experiments results

Algorithm	Acc. train envs.	Acc. test env.
ERM IRM (ours)	$87.4 \pm 0.2 \\ 70.8 \pm 0.9$	$\begin{array}{c} 17.1\pm0.6\\ \textbf{66.9}\pm\textbf{2.5} \end{array}$
Random guessing (hypothetical) Optimal invariant model (hypothetical) ERM, grayscale model (oracle)	$50 \\ 75 \\ 73.5 \pm 0.2$	$50 \\ 75 \\ 73.0 \pm 0.4$

Table 1: Accuracy (%) of different algorithms on the Colored MNIST synthetic task. ERM fails in the test environment because it relies on spurious color correlations to classify digits. IRM detects that the color has a spurious correlation with the label and thus uses only the digit to predict, obtaining better generalization to the new unseen test environment.



Figure 5: P(y = 1|h) as a function of h for different models trained on Colored MNIST: (left) an ERM-trained model, (center) an IRM-trained model, and (right) an ERM-trained model which only sees grayscale images and therefore is perfectly invariant by construction. IRM learns approximate invariance from data alone and generalizes well to the test environment.

Authors: Martin Arjovsky, L´eon Bottou, Isha

Invariant Risk Minimization